Multi-View Regression via Canonincal Correlation Analysis

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(Sham M. Kakade of TTI)

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$$(\forall i \leq n)$$
 $y_i = \sum_{i=j}^{p} X_{ij}\beta_j + \epsilon_i \quad \epsilon_i \sim_{iid} N(0, \sigma^2)$

Data mining and Machine learning: $p \gg n$



$$(\forall i \leq n)$$
 $y_i = \sum_{i=j}^{p} X_{ij}\beta_j + \epsilon_i \quad \epsilon_i \sim_{iid} N(0, \sigma^2)$

Can't fit model if $p \gg n$:

- Trick: assume most β_i are in fact zero
- Variable selection:

$$\hat{\beta}_{i}^{\mathsf{RIC}} = \begin{cases} 0 & \text{if } |\hat{\beta}_{i}| \leq SE_{i}\sqrt{2\log p} \\ \hat{\beta}_{i} & \text{otherwise} \end{cases}$$

- Basically just stepwise regression and Bonferroni
 - Can be justified by "risk ratios" (Donoho and Johnstone '94, Foster and George '94)

$$(\forall i \leq n)$$
 $y_i = \sum_{i=j}^{p} X_{ij}\beta_j + \epsilon_i \quad \epsilon_i \sim_{iid} N(0, \sigma^2)$

I've played with lots of alternatives:

- FDR instead of RIC:
 - $\sqrt{2\log p} \rightarrow \sqrt{2\log(p/q)}$
 - empirical Bayes (George and Foster, 2000)
 - Cauchy prior (Foster and Stine, 200x)
- $\bullet \ regression \rightarrow logistic \ regression$
- IID \rightarrow independence
- independence → block independence (with Dongyu Lin)

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$$(\forall i \leq n)$$
 $y_i = \sum_{i=j}^{p} X_{ij}\beta_j + \epsilon_i \quad \epsilon_i \sim_{iid} N(0, \sigma^2)$

Where do this many variables come from?

- Missing value codes
- Interactions
- Transformations
- Historical example (Personal Bankruptcy)
 - 350 basic variables
 - all interactions, missing value codes, etc lead to 67,000 variables
 - about 1 million clustered cases
 - Ran stepwise logistic regression using FDR
 - Another talk tells of the details of that experiment

$$(\forall i \leq n)$$
 $y_i = \sum_{i=j}^{p} X_{ij}\beta_j + \epsilon_i \quad \epsilon_i \sim_{iid} N(0, \sigma^2)$

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Summary of current state of the art:

- We can generate many non-linear X's
- We can select the good ones large lists
- Isn't the problem "solved"?

$$(\forall i \leq n)$$
 $y_i = \sum_{i=j}^{p} X_{ij}\beta_j + \epsilon_i \quad \epsilon_i \sim_{iid} N(0, \sigma^2)$

There is always room for finding new X's



- Current methods of finding X's are non-linear
- Can we find "new" linear combinations of existing X's?

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- Hope, use linear theory
- Hope, fast CPU
- Hope, new theory

Semi-supervised learning is:

- Y's are expensive
- X's are cheap
- We get *n* rows of *Y*
- But also *m* free rows of just X's
- Called, semi-supervised learning

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Can this help?

Usual data table for data mining

$$\begin{bmatrix} Y \\ (n \times 1) \end{bmatrix} \begin{bmatrix} X \\ (n \times p) \end{bmatrix}$$

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with $p \gg n$

With unlabeled data

m rows of unlabeled data:

$$\left[\begin{array}{c} Y\\ n\times 1 \end{array}\right] \qquad \left[\begin{array}{c} X\\ (n+m)\times p \end{array}\right]$$

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m rows of unlabeled data, and two sets of equally useful *X*'s:

$$\begin{bmatrix} Y \\ n \times 1 \end{bmatrix} \begin{bmatrix} X \\ (n+m) \times p \end{bmatrix} \begin{bmatrix} Z \\ (n+m) \times p \end{bmatrix}$$

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With: $m \gg n$

Examples

- Person identification
 - Y = identity
 - X = Profile photo
 - Z = front photo
- Topic identification (medline)
 - Y = topic
 - X = abstract
 - Z = text
- The web:
 - Y = classification
 - X = content (i.e. words)
 - Z = hyper-links
- We will call these the multi-view setup

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A Multi-View Assumption

Define

$$\sigma_x^2 = E[Y - E(Y|X)]^2$$

$$\sigma_z^2 = E[Y - E(Y|Z)]^2$$

$$\sigma_{x,z}^2 = E[Y - E(Y|X,Z)]^2$$

(We will take conditional expectations to be linear)

Assumption

Y,X, and Z satisfy the α -multiview assumption if:

$$\begin{array}{rcl} \sigma_x^2 &\leq & \sigma_{x,z}^2(1+\alpha) \\ \sigma_z^2 &\leq & \sigma_{x,z}^2(1+\alpha) \end{array}$$

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• In other words, $\sigma_x^2 \approx \sigma_z^2 \approx \sigma_{x,z}^2$

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• Views X and Z are redundant (i.e. highly collinear)

The Multi-View Assumption in the Linear Case

- The views are redundant.
- Satisfied if each view predict Y well.
- No conditional independence assumptions (i.e. Bayes nets)
- No coordinates, norm, eigenvalues, or dimensionality assumptions.

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Lemma

Under the α -multiview assumption

$$E[(E(Y|X) - E(Y|Z))^2] \le 2\alpha\sigma^2$$

• Idea: find directions in X and Z that are highly correlated

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• CCA solves this problem already!

What if we run CCA on X and Z?

CCA = canonical correlation analysis

- Find the directions that are most highly correlated
- Very close to PCA (principal components analysis)
- Generates coordinates for data
- End up with canonical coordinates for both X's and Z's

• Numerically an Eigen-value problem

Definition

 X_i , and Z_j , are in CCA form if

- X_i are orthonormal
- *Z_i* are orthonormal

•
$$X_i^T Z_j = 0$$
 for $i \neq j$

•
$$X_i^T Z_i = \lambda_i, \, \lambda_1 \geq \lambda_2 \geq \cdots \geq \lambda_p \geq 0$$

(This is the output of running CCA on the original X's and Z's.)

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CCA form as a covariance matrix

$$\Sigma = \begin{bmatrix} \underline{\Sigma_{XX}} & \underline{\Sigma_{XZ}} \\ \underline{\Sigma_{ZX}} & \underline{\Sigma_{ZZ}} \end{bmatrix} \rightarrow \begin{bmatrix} I & D \\ \hline D & I \end{bmatrix}$$

The *canonical correlations* are λ_i :

$$D = \begin{bmatrix} \lambda_1 & 0 & 0 & \dots \\ 0 & \lambda_2 & 0 & \dots \\ 0 & 0 & \lambda_3 & \dots \\ \vdots & \vdots & \vdots & \vdots \end{bmatrix}$$

Let $\hat{\beta}$ be the Ridge regression estimator with weights induced by the CCA. Then

$$\textit{\textit{Risk}}(\hat{eta}) \leq \left(5lpha + rac{\sum \lambda_i^2}{n}
ight)\sigma^2$$

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Let $\hat{\beta}$ be the Ridge regression estimator with weights induced by the CCA. Then

$$extsf{Risk}(\hat{eta}) \leq \left(5lpha + rac{\sum \lambda_i^2}{n}
ight) \sigma^2$$

CCA-ridge regression is to minimize least squares plus a penalty of:

$$\sum_{i} \frac{1 - \lambda_i}{\lambda_i} \beta_i^2$$

- Large penalties in the less correlated directions.
- λ_i 's are the correlations
- A shrinkage estimator.

Let $\hat{\beta}$ be the Ridge regression estimator with weights induced by the CCA. Then

$$extsf{Risk}(\hat{eta}) \leq \left(5lpha + rac{\sum \lambda_i^2}{n}
ight) \sigma^2$$

Recall α is the multiview property:

$$\sigma_x^2 \leq \sigma_{x,z}^2(1+\alpha)$$

$$\sigma_z^2 \leq \sigma_{x,z}^2(1+\alpha)$$

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• 5α is the bias • $\frac{\sum \lambda_i^2}{n}$ is variance

Let $\hat{\beta}$ be the Ridge regression estimator with weights induced by the CCA. Then

$$\textit{\textit{Risk}}(\hat{eta}) \leq \left(5lpha + rac{\sum \lambda_i^2}{n}
ight)\sigma^2$$

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Doesn't fit my personality and style

- I like feature selection!
- On to theorem 2

Alternative version

Theorem

For $\hat{\beta}$ be the CCA-testimator:

$$\mathsf{Risk}(\hat{eta}) \leq \left(2\sqrt{lpha} + rac{\mathsf{d}}{\mathsf{n}}
ight)\sigma^2$$

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where d is the number of λ_i for which $\lambda_i \geq 1 - \sqrt{\alpha}$.

For $\hat{\beta}$ be the CCA-testimator:

$$\mathsf{Risk}(\hat{eta}) \leq \left(2\sqrt{lpha} + rac{\mathsf{d}}{\mathsf{n}}
ight) \sigma^2$$

where d is the number of λ_i for which $\lambda_i \geq 1 - \sqrt{\alpha}$.

The CCA testimator:

$$\widehat{\beta}_{i} = \begin{cases} \text{MLE}(\beta_{i}) & \text{if } \lambda_{i} \ge 1 - \sqrt{\alpha} \\ 0 & \text{else} \end{cases}$$
(1)

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For $\hat{\beta}$ be the CCA-testimator:

$$\mathsf{Risk}(\hat{eta}) \leq \left(2\sqrt{\alpha} + \frac{\mathsf{d}}{\mathsf{n}}\right)\sigma^2$$

where d is the number of λ_i for which $\lambda_i \geq 1 - \sqrt{\alpha}$.

Do we need to know α ?

- We can try features in order
- Use promiscuous rule to add variables (i.e. AIC)
- Will do as well as theorem, and possibly much better
- Doesn't mix all that well with stepwise regression

• Trade off between two theorems?

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Experimental work?

• Trade off between two theorems?

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Experimental work? Soon!